# Analysis of AHP by BIBD 

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Abstract The essence of AHP is to evaluate objects in terms of the eigen vector of the comparison matrix. But when the number of objects, $n$, is too large, it causes often worse reliability for an observer to evaluate all paired comparisons at a time. So it is necessary to decompose the whole set of pairs into several classes, and for each class to be evaluated by one observer. We propose the decomposition by BIBD (balanced incomplete block design) well known in the field of experimental design or combinatorics. We show by simulation experiments that our method gives better evaluations than the ordinary AHP.

In connection with these, we show that the logarithmic least square method, which has been proposed by several authors, gives very good approximation to the eigen vector method when $n$ is rather small, and that the former completely coincides with the latter when $n \leq 3$, surprisingly.

## 1. Introduction

The essence of AHP (Analytic Hierarchy Process) is to evaluate objects in terms of the eigen vector corresponding to the maximal eigen value of matrix whose $(i, j)$ element is the ratio of evaluation of object $i$ to object $j[1,2,3]$. The idea is to intend to unify local informations taken by paired comparison into a global information.

But when the number of objects is too large, it causes often worse reliability for a observer to evaluate all paired comparisons at a time. In such case, it is necessary to decompose the whole set of pairs of objects into several blocks, and for each block to be observed by one observer. It is important how to decompose the set of pairs. We propose the decomposition by BIBD (Balanced Incomplete Block Design) well known in the field of experimental design $[4,5]$. And we show that this method is to give better evaluation under certain assumptions by simulation experiments (§4). Further we propose the logarithmic least square method for our problem, and show that this gives a good approximation to the eigen value method in AHP (§3).

## 2. BIBD

Let $E=\{1,2, \cdots, v\}$ be the set of objects $i=1,2, \cdots, v$. The number of pairs among $E$ is ${ }_{v} C_{2}=v(v-1) / 2$, and if $v$ is large this becomes very large and an observer cannot compare all such pairs at a time. Let the set of objects which an observer can accommodate with sufficient reliability, be called an "allowable block", and let us denote the size of allowable block by $k(\leq v)$.

Then we need to decompose the set of ${ }_{v} C_{2}$ pairs into the classes of size ${ }_{k} C_{2}$ and to allocate several observers to these classes. Each observer makes paired comparisons in his class. We unify these results and can get the evaluation on $E$.

For example, there are $v=7$ applicants for a prize essay contest, and we try to judge their essays and to decide ranking on them. Let the allowable block size be $k=3$, that is, one judge can read 3 essays and make paired comparisons on them. In this case, the ${ }_{7} C_{2}=21$ pairs are decomposed into classes of size ${ }_{3} C_{2}=3$. So we need $21 / 3=7$ judges.

We unify the results of 7 judges into the whole ranking on 7 applicants.
It is the problem how to decompose the set of ${ }_{"} C_{2}$ pairs into blocks and how to unify the results of observations on blocks into the whole evaluation. We propose the decomposition by BIBD and the unification by the eigen value analysis used in AHP.

BIBD on $E=\{1,2, \cdots, v\}$ is the class $\mathbf{D}=\left\{B_{1}, \cdots, B_{b}\right\}$ of subsets (called "blocks") $B_{e} \subseteq E(e=1, \cdots, b)$ satisfying the followings; (i) for any $t=1, \cdots, b$
$\left|B_{e}\right|=k$ (the block size is constant)
(ii) for any $i=1, \cdots, v$
$\left|\left\{e \mid i \in B_{e}, e=1, \cdots, b\right\}\right|=r$ (the replication number is constant)
(iii) for any $i, j=1, \cdots, v(i \neq j)$
$\left|\left\{e \mid i \in B_{e}, j \in B_{e}, e=1, \cdots, b\right\}\right|=\lambda$ (the intersection number is constant)
(where $|S|$ denotes the size of a set $S$.)
It is clear that (iii) implies (ii), so (i) and (iii) suffice for $\mathbf{D}$ to be BIBD. Specifically a BIBD with $\lambda=1$ is called Steiner system and here we consider only Steiner system, which is denoted by $(v, k)-\mathbf{D}$.

Example 1. The class of subsets of $E=\{1,2, \cdots, 7\}$ shown in Table 1 is (7, 3)-D. The pairs in each block are shown on the right hand of the block. All such pairs construct the set of ${ }_{7} C_{2}$ pairs. In other word, the set of pairs in $E$ is decomposed into pairs in blocks.

We can represent this situation in terms of graph theory in the following way; To construct $(v, k)$ - $\mathbf{D}$ is equivalent to decompose the set of edges complete graph with $v$ points into $b$ complete graphs each of which has $k$ points. ( $\rightarrow$ Fig. 1)

Table 1 (7, 3).D

$$
\begin{aligned}
& B_{1}=\{1,2,4\} \rightarrow 12,14,24 \\
& B_{2}=\{2,3,5\} \rightarrow 23,25,35 \\
& B_{3}=\{3,4,6\} \rightarrow 34,36,46 \\
& B_{4}=\{4,5,7\} \rightarrow 45,47,57 \\
& B_{5}=\{5,6,1\} \rightarrow 56,51,61 \\
& B_{6}=\{6,7,2\} \rightarrow 67,62,72 \\
& B_{7}=\{7,1,3\} \rightarrow 71,73,13
\end{aligned}
$$



Fig. 1 BIBD decomposition

By the graph theoretic representation stated in Example 1 we can easily have the following relations.

$$
\begin{gather*}
v r=b k  \tag{2.1}\\
b=\frac{{ }_{v} C_{2}}{{ }_{k} C_{2}}=\frac{v(v-1)}{k(k-1)} \tag{2.2}
\end{gather*}
$$

If a $(v, k)$-D exists then integers $v$ and $k$ satisfy (2.1) and (2.2) with integers $r$ and $b$, so for any integers $v$ and $k$ we do not necessarily have $(v, k)$-D. For example ( 8,3 )-D never exists. But (9, 3)-D exists, so we can treat the case $v=8, k=3$ by taking one of objects in ( 9,3 )- $\mathbf{D}$ as a dummy.

The conditions of existence and construction methods of $(v, k)$ - $\mathbf{D}$ have been widely and deeply researched in the field of experimental designs and combinatorial theories $[4,5]$.

In order to show why the decomposition by the $(v, k)$ - $\mathbf{D}$ is appropriate, we propose another rather natural decomposition shown in Table 2. Of course this $\mathbf{D}=\left\{B_{1}, \cdots B_{7}\right\}$ is not BIBD, where pair $(1,2)$ occurs twice in $B_{1}$ and $B_{7}$, while pairs $(1,4)$ and $(1,5)$ do not occur anywhere.

The author believes that $(v, k)$-D would give the best possible decompositions for our problems. At the end of this section we give another $(v, k)$ - $\mathbf{D}$ in Table 3.

## Table $2(7,3) \cdot \mathrm{D}$

$$
\begin{aligned}
& B_{1}=\{1,2,3\} \rightarrow 12,13,23 \\
& B_{2}=\{2,3,4\} \rightarrow 23,24,34 \\
& B_{3}=\{3,4,5\} \rightarrow 34,35,45 \\
& B_{4}=\{4,5,6\} \rightarrow 45,46,56 \\
& B_{5}=\{5,6,7\} \rightarrow 56,57,67 \\
& B_{6}=\{6,7,1\} \rightarrow 67,61,71 \\
& B_{7}=\{7,1,2\} \rightarrow 71,72,12
\end{aligned}
$$

Table 3 (13, 4)-D


## 3. LLS, AHP estimation

Let $\{1, \cdots, n\}$ be the set of objects $i=1,2, \cdots, n$. An observer observes the ratio of evaluation object $i$ to object $j$ and let us denote the observation by $x_{i j}$.

We assume the statistical model of $x_{i j}$ to be

$$
\begin{align*}
& x_{i j}=a_{i j} \cdot e_{i j}, \quad a_{i j}=w_{i} / w_{j} \quad(i<j, i, j=1 \sim n) \\
& x_{j i}=1 / x_{i j} \tag{3.1}
\end{align*}
$$

Where $w_{i}(>0)$ is the real evaluation of object $i$ and is an unknown parameter, and $e_{i j}(>0)$ is an independent random variable representing the error of the observation. And we always recognize that any multiples of $\left[w_{1} \cdots w_{n}\right]$ are equivalent to $\left[w_{1} \cdots w_{n}\right]$ itself. Further we assume that

$$
\begin{equation*}
E\left(\ln e_{i j}\right)=0, V\left(\ln e_{i j}\right)=\sigma^{2}(n) \tag{3.2}
\end{equation*}
$$

and $\sigma^{2}(n)$ is a monotone increasing function of $n$, the number of objects to be observed.
Whether these assumptions are reasonable or not is a psychological or a physiological problem, but we can agree with these assumptions as a trial scheme.

The main purpose of AHP is to get estimates $\hat{w}_{i}$ of $w_{i}(i=1, \cdots, n)$ by calculating the eigen vector $\hat{w}=\left[\hat{w}_{1}, \cdots, \hat{w}_{n}\right]$ corresponding to the maximal eigen value $\lambda$ of the $n \times n$ comparison matrix $X=\left[x_{i j}\right]$.

Of course we have other estimation methods. The most natural one is "logarithmic least square (LLS)". For simplicity let $\bar{x}_{i j}=\ln x_{i j}, \bar{w}_{i}=\ln w_{i}$ and $\bar{e}_{i j}=\ln e_{i j}$. Then we have

$$
\begin{equation*}
\bar{x}_{i j}=\bar{w}_{i}-\bar{w}_{j}+\bar{e}_{i j}(i<i, i, j=1, \cdots, n) \tag{3.3}
\end{equation*}
$$

Appling the least square method to (3.3) we have least square estimate $\hat{\bar{w}}_{i}$ of $\bar{w}_{i}$, and taking inverse transform we have $\hat{w}_{i}=\exp \left(\hat{\bar{w}}_{i}\right)(i==1, \cdots, n)$. This is LLS estimation.

For example let $n=3$, then we have

$$
\begin{equation*}
\bar{x}_{12}=\bar{w}_{1}-\bar{w}_{2}+\bar{e}_{12}, \bar{x}_{13}=\bar{w}_{1}-\bar{w}_{3}+\bar{e}_{13}, \bar{x}_{23}=\bar{w}_{2}-\bar{w}_{3}+\bar{e}_{23} \tag{3.4}
\end{equation*}
$$

As the vector $w=\left[w_{1}, \cdots, w_{n}\right]$ multiplied by an arbitrary constant is equivalent to $w$ itself, we can assume $w_{1} w_{2} w_{3}=1$, so we have

$$
\begin{equation*}
\bar{w}_{1}+\bar{w}_{2}+\bar{w}_{3}=0 \tag{3.5}
\end{equation*}
$$

Applying least square method to (3.4), (3.5) and taking inverse transform we have

$$
\begin{equation*}
\hat{w}_{1}=\left(x_{12} x_{13}\right)^{1 / 3}, \hat{w}_{2}=\left(x_{23} x_{21}\right)^{1 / 3}, \hat{w}_{3}=\left(x_{31} x_{32}\right)^{1 / 3} \tag{3.6}
\end{equation*}
$$

This estimation is very simple, but we have a surprising fact that $w_{i}(i=1,2,3)$ in (3.6) coincide with the components of the eigenvector corresponding to the maximal eigenvalue of the comparison matrix $X=\left[x_{i j}\right]$. That is, we have

Theorem 1. In the case $\leq 3$, LLS estimates coincide with AHP estimates.
Proof: In case of $n=2$, we have easily

$$
\begin{equation*}
\hat{w}_{1}=\sqrt{x_{12}}, \hat{w}_{2}=\sqrt{x_{21}} \tag{3.7}
\end{equation*}
$$

as the LLS estimates of $w_{1}, w_{2}$ respectively. And by direct calculation we have

$$
X\left[\begin{array}{c}
\hat{w}_{1} \\
\hat{w}_{2}
\end{array}\right]=2\left[\begin{array}{c}
\hat{w}_{1} \\
\hat{w}_{2}
\end{array}\right] \quad X=\left[\begin{array}{cc}
1 & x_{12} \\
x_{21} & 1
\end{array}\right]
$$

and the maximal eigen value of $2 \times 2$ comparison matrix is 2 , so $\hat{w}_{1} \hat{w}_{2}$ in (3.7) are the AHP estimates.

In case of $n=3$, from (3.6) we have

$$
\begin{aligned}
& {\left[\begin{array}{ccc}
1 & x_{12} & x_{13} \\
x_{21} & 1 & x_{23} \\
x_{31} & x_{32} & 1
\end{array}\right]\left[\begin{array}{l}
\hat{w}_{1} \\
w_{2} \\
\hat{w}_{3}
\end{array}\right]=\lambda\left[\begin{array}{l}
\hat{w}_{1} \\
\hat{w}_{2} \\
\hat{w}_{3}
\end{array}\right]} \\
& \lambda=1+r+1 / r, r==\sqrt[3]{x_{12} x_{23} x_{31}}
\end{aligned}
$$

by direct calculation. But from Perron \& Frobenius Theorem we can state that if a positive matrix has a positive eigen vector then this corresponds to the maximal eigen value, ( A positive matrix (vector) means a matrix (vector) whose components are all positive). (See the proof of the theorem 4 in [6]).

Like the case $n=3$ we have general LLS estimates

$$
\begin{equation*}
\hat{w}_{i}=\left(\Pi_{j=1}^{n} x_{i j}\right)^{1 / n} ;=1, \cdots, n \tag{3.8}
\end{equation*}
$$

which can be stated simply as $\hat{w}_{i}$ is the geometric mean of the $i$-th row of the comparison matrix $X$.

But unfortunately they no longer coincide with AHP estimates for $n>3$. Thus if $n>3$ then Theorem 1 does not hold. But through Theorem 1 we can state that if $\sigma^{2}(n)$ is reasonably small LLS estimates must be good approximations to AHP estimates even if $n$ is greater than 3. This also teaches us that our statistical model (3.1) is valid for the AHP analysis.

## 4. Decomposition Methods by BIBD

Now we propose our decomposition methods by BIBD. We are given the set of objects $E=\{1,2, \cdots, v\}$ to be evaluated. Let the allowable block size $k$ be far smaller than $v$. We decompose $E$ into blocks $B_{1}, \cdots, B_{b}$ which construct Steiner system $(v, k)$-D.

Step 1. For each $B_{e}=\left\{\beta_{1}, \beta_{2}, \cdots, \beta_{k}\right\} \subseteq E$ an observer observes the objects and gets observation $x_{\beta_{t} \beta}$, by a paired comparison $\beta_{t}$ to $\beta_{s}(t<x ; t, s=1, \cdots, k)$. Note that the observation error of $x_{\beta_{t} \beta_{s}}$ (measured by $V\left(\ln x_{\beta_{\varepsilon} \beta_{s}}=\sigma^{2}(k)\right)$ is far smaller than the one incurred by the observation in the whole set $E$.
Step 2. Let

$$
\begin{equation*}
S_{e}=\left\{x_{\beta_{1}, \beta_{2}}, x_{\beta_{1} \beta_{3}}, \cdots, x_{\beta_{k-1} \beta_{k}}\right\} \tag{4.1}
\end{equation*}
$$

and let

$$
\begin{aligned}
\bar{S}_{e} & =\left\{x_{\beta_{2} \beta_{1}}, x_{\beta_{3} \beta_{1}}, \cdots, x_{\beta_{k} \beta_{k-1}}\right\} \\
& =\left\{1 / x_{\beta_{1} \beta_{2}}, 1 / x_{\beta_{1} \beta_{3}}, \cdots, 1 / x_{\beta_{k-1} \beta_{k}}\right\}
\end{aligned}
$$

$$
\begin{equation*}
\text { for } e=1,2, \cdots b \text {. } \tag{4.2}
\end{equation*}
$$

Then by the properties of BIBD

$$
\left(S_{1}^{U} \bar{S}_{1}\right)^{U}\left(S_{2}^{U} \bar{S}_{2}\right)^{U} \ldots{ }^{U}\left(S_{b}^{U} \bar{S}_{b}\right)
$$

constructs the set of all paired comparisons $x_{i j}(i, j=1, \cdots, v)$ from which we construct the $v \times v$ comparison matrix $X=\left[x_{i j}\right]$
Step 3. We apply the usual AHP to the comparison matirix $X$, that is, we calculate the eigen vector $\mathbf{w}=\left[\hat{w}_{1}, \hat{w}_{2}, \cdots, \hat{w}_{v}\right]$ corresponding the maximal eigen value of $X$. Then $\hat{w}_{i}$ is the desired estimate of $w_{i}$ in (3.1) $(i=1, \cdots, n)$.
We will show by several simulation examples that this decomposition method has higher reliability than the ordinary direct method.

Example 3. ( $v=7, k=3$, BIBD decomposition)
Let $w_{1}, \cdots, w_{7}$ of (3.1) take the values shown in Table 4. Of course these are unknown for the observers and are to be estimated. Further $a_{i j}=w_{i} / w_{j}(i, j=1, \cdots, 7)$ are also shown in Table 4.

For each block $B_{e}$ (shown in Table 2) an observer takes observations $x_{\beta_{1}, \beta_{2}}, x_{\beta_{1} \beta_{3}}, x_{\beta_{2} \beta_{3}}$, where

$$
\begin{equation*}
x_{\beta_{t} \beta_{s}}=a_{\beta_{t} \beta_{s}} e_{\beta_{t} \beta_{s}} \tag{4.3}
\end{equation*}
$$

and $e_{\beta_{\mathrm{t}} \beta_{s}}$ is a random number whose logarithm $\ln e_{\beta_{\mathrm{t}} \beta_{s}}$ normally distributes with zero mean and variance $\sigma^{2}(3)$, by our model, (3.1) and (3.2) (For the actual value of $\sigma^{2}(3)$, see (4.5)). For $e=1,2, \cdots, 7 S_{e}=\left\{x_{\beta_{1} \beta_{2}}, x_{\beta_{1} \beta_{3}}, x_{\beta_{2} \beta_{3}}\right\}$ and $\bar{S}_{e}=\left\{x_{\beta_{2} \beta_{1}}=1 / x_{\beta_{3} \beta_{1}}=1 / x_{\beta_{1} \beta_{3}}, x_{\beta_{3} \beta_{2}}=\right.$ $\left.1 / x_{\beta_{2} \beta_{3}}\right\}$ are whon in Table 5 in the matrix form

$$
X_{e}=\left[\begin{array}{ccc}
1 & x_{\beta_{1} \beta_{2}} & x_{\beta_{1} \beta_{3}}  \tag{4.4}\\
x_{\beta_{2} \beta_{1}} & 1 & x_{\beta_{2}} \\
x_{\beta_{3} \beta_{1}} & x_{\beta_{3} \beta_{2}} & 1
\end{array}\right]
$$

Unifying $X_{1}, \cdots, X_{7}$ we have the $7 \times 7$ comparison matrix $X$ shown in Table 6 , and calculating the eigen vector for the maximal eigen value $\lambda$ of $X$ we have the estimates $\hat{w}_{1}, \cdots, \hat{w}_{7}$ shown also in Table 6. Comparing $\hat{w}_{i}$ to $w_{i}(i=1 \sim 7)$ we can say that we have generally fairly good estimates.


Table $6 \quad 7 \times 7 \quad X=\left[x_{i j}\right]$ (BIBD decomposition)

|  | 1 | 2 | 3 | 4 | 5 | 6 | 7 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
|  | 1 | 1 | 1.540 | 19.33 | 7.236 | 8.314 | 3.800 |
| 2 | 0.6493 | 1 | 13.046 | 4.269 | 6.171 | 1.881 | 2.760 |
| 3 | 0.05173 | 0.07665 | 1 | 0.4033 | 0.4057 | 0.1975 | 0.2293 |
| 4 | 0.1382 | 0.2342 | 2.479 | 1 | 1.1930 | 0.1129 | 0.5590 |
| 5 | 0.1135 | 0.1621 | 2.465 | 0.9709 | 1 | 0.3887 | 0.5544 |
| 6 | 0.2632 | 0.5317 | 5.063 | 3.004 | 2.572 | 1 | 1.226 |
| 7 | 0.1870 | 0.3623 | 4.361 | 1.789 | 1.304 | 0.8160 | 1 |

$$
\begin{aligned}
& \lambda=7.0294 \\
& \dot{w}_{1}=1 \\
& w_{2}=0.6067 \\
& \dot{w}_{3}=0.0498 \\
& \mathfrak{w}_{4}=0.1222 \\
& w_{5}=0.1136 \\
& \dot{w}_{6}=0.2916 \\
& \dot{w}_{7}=0.2132
\end{aligned}
$$

Now we consider the actual value of $\sigma^{2}(n)$ in our model (3.1). Of course this depends on the given real problem. But we assume

$$
\begin{equation*}
\sigma(3)^{2}=0.158^{2}, \sigma^{2}(4)=0.250^{2}, \sigma^{2}(7)=0.5^{2}, \sigma^{2}(13)=0.980^{2} \tag{4.5}
\end{equation*}
$$

as trial values in our simulations.
Here we assume that $\sigma^{2}=\sigma^{2}(n)$ is a linear function of the number $N={ }_{n} C_{2}$ of paired comparisons, that is $\sigma^{2}=a+b N$. Of course this is not based on any psychological or physiological knowledges, but it is certain that $\sigma^{2}$ depends on $N$ and it is rather natural to assume a linear function as a first approximation. Now $a$ is determined by the condition; $\sigma^{2}=0$ when $N=1$. And $b$ is to be arbitrary as a scale parameter. Here we select the value of $b$ by the condition $\sigma^{2}(7)=0.5^{2}$ which is a standard case.

Example 4. ( $v=7$, direct method)
Here we will describe the usual AHP method applied directly to Table 4. We multiply $a_{i j}$ by a random number $e_{i j}$ and have

$$
x_{i j}=a_{i j} e_{i j}(i<j, i, j=1, \cdots, 7)
$$

where $\ln e_{i j}$ normally distributes with zero mean and variance $\sigma^{2}(7)=0.5^{2}(\rightarrow(4.5))$, and calculate $x_{j i}=1 / x_{i j}$ then we have the comparison matrix $X=\left[x_{i j}\right]$ shown in Table 7 . Calculating the eigen vector for the maximal eigen value $\lambda$ we have estimates $\hat{w}_{i}(i=1, \cdots, 7)$ shown also in Table 7, quite worse than delete the ones in Table 6.

Table $7 \quad X=\left[x_{i j}\right]$ (usual AHP)

| 1 | 2 | 3 | 4 | 5 | 6 | 7 | $\lambda=7.323$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.731 | 13.29 | 7.942 | 8.944 | 1.437 | 4.561 | $\hat{w}_{1}=1$ |
| 0.5777 | 1 | 11.68 | 4.720 | 6.245 | 2.188 | 4.173 | $\hat{w}_{2}=0.7679$ |
| 0.07524 | 0.08562 | 1 | 0.3694 | 1.190 | 0.1597 | 0.3299 | $\hat{w}_{3}=0.0793$ |
| 0.1259 | 0.2119 | 0.707 | 1 | 0.5543 | 0.3720 | 0.3892 | $\hat{w}_{4}=0.1303$ |
| 0.1118 | 0.1601 | 0.8402 | 1.804 | 1 | 0.5835 | 0.6611 | $\dot{w}_{5}=0.1460$ |
| 0.6959 | 0.4570 | 6.262 | 2.688 | 1.714 | 1 | 2.150 | $\hat{w}_{6}=0.4139$ |
| 0.2193 | 0.2396 | 3.031 | 2.569 | 1.513 | 0.4651 | 1 | $\hat{w}_{7}=0.2201$ |

Next we try to investigate another case $v=13, k=4$ in Examples 5, 6 along the same line as above.

Example 5. ( $v=13, k=4$, BIBD decomposition)
We show $w_{i}(i=1, \cdots, 13)$ and $a_{i j}=w_{i} / w_{j}(i, j=1, \cdots, 13)$ in Table 8, and (13,4)-D and its comparison observations in Table 9, where the logarithm of the random numbers have the variance $\sigma^{4}(4)=0.250^{2}$ (in (4.5)). Finally the unified comparison matrix $X$ and its maximal eigen value and the corresponding eigen vector are shown in Table 10. These give us rather good estimates.

Example 6. ( $v=13$, direct method)
If an observer observes ${ }_{13} C_{2}$ paired comparisons $x_{i j}(i<j ; i, j=1 \sim 13)$ at a time, then the logarithm of observation error $e_{i j}$ has variance $\sigma^{2}(13)=0.980^{2}(\rightarrow(4.5))$. We calculate $x_{i j}=a_{i j} e_{i j}$ from Table 8 and the random numbers $e_{i j}$ with above mentioned properties. $X=\left[x_{i j}\right.$ and its maximal eigen value and its eigen vector are shown in Table 11. These estimates almost have no reliability.

|  |  | 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| $w_{1}=0.1341$ | 1 | 1 | 1.111 | 1.235 | 1.372 | 1.524 | 1.6393 | 1.882 | 2.091 | 2.323 | 2.581 | 2.868 | 3.187 | 3.541 |
| $w_{2}=0.1207$ | 2 | 0.9 | 1 | 1.111 | 1.235 | 1.372 | 1.524 | 1.693 | 1.882 | 2.091 | 2.323 | 2.581 | 2.868 | 3.187 |
| $w_{3}=0.1086$ | 3 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.382 | 1.524 | 1.693 | 1.882 | 2.091 | 2.323 | 2.581 | 2.868 |
| $w_{4}=0.0977$ | 4 | 0.729 | 0.81 | 0.9 | 1 | 1.235 | 1.372 | 1.524 | 1.693 | 1.882 | 2.091 | 2.323 | 2.581 |  |
| $w_{5}=0.0880$ | 5 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.372 | 1.524 | 1.693 | 1.882 | 2.091 | 2.323 |
| $w_{6}=0.0792$ | 6 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.372 | 1.524 | 1.693 | 1.882 | 2.091 |
| $\mathrm{w}_{7}=0.0713$ | 7 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.372 | 1.524 | 1.693 | 1.882 |
| $\mathbf{w}_{\text {B }}=0.0641$ | 8 | 0.4783 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.372 | 1.524 | 1.693 |
| $w_{9}=0.0577$ | 9 | 0.4305 | 0.4783 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.372 | 1.524 |
| $\mathbf{w}_{10}=0.0519$ | 10 | 0.3874 | 0.4305 | 0.4783 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 | 1.235 | 1.372 |
| $w_{11}=0.0468$ | 11 | 0.3487 | 0.3874 | 0.4305 | 0.4783 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | $0.9{ }^{\circ}$ | 1 | 1.111 | 1.235 |
| $w_{12}=0.0421$ | 12 | 0.3138 | 0.3487 | 0.3874 | 0.4305 | 0.4783 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 | 1.111 |
| $\mathbf{w}_{13}=0.0379$ | 13 | 0.2824 | 0.3138 | 0.3487 | 0.3874 | 0.4305 | 0.4783 | 0.5314 | 0.5905 | 0.6561 | 0.729 | 0.81 | 0.9 | 1 |



Table $10 \quad X_{1}, X_{2}, \ldots . X_{13}$

| 1.0000 | 1.2720 | 1.7600 | 1.0570 | 0.9139 | 1.7120 | 2.2150 | 1.9590 | 1.8650 | 3.1440 | 4.4770 | 2.1470 | 4.9400 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 0.7862 | 1.0000 | 0.9462 | 0.9282 | 2.1130 | 1.3980 | 1.7440 | 1.6910 | 2.6970 | 1.4150 | 3.0660 | 3.7910 | 4.0540 |
| 0.5682 | 1.0570 | 1.0000 | 1.1670 | 1.3880 | 1.0550 | 1.8670 | 1.3240 | 1.9700 | 1.3850 | 1.6660 | 2.4070 | 2.9660 |
| 0.9461 | 1.0770 | 0.8569 | 1.0000 | 1.1890 | 0.8383 | 1.5500 | 1.0420 | 1.7320 | 2.6480 | 2.0870 | 2.6670 | 1.5230 |
| 1.0940 | 0.4733 | 0.7205 | 0.8410 | 1.0000 | 1.3280 | 0.7647 | 1.5230 | 1.3660 | 1.6130 | 1.9370 | 1.7690 | 1.9490 |
| 0.5841 | 0.7153 | 0.9479 | 1.1930 | 0.7530 | 1.0000 | 1.1350 | 1.1470 | 1.9040 | 1.1620 | 2.5040 | 1.9540 | 2.9150 |
| 0.4515 | 0.5734 | 0.5356 | 0.6452 | 1.3080 | 0.8811 | 1.0000 | 1.0770 | 1.1500 | 0.9667 | 1.9450 | 1.6020 | 1.8460 |
| 0.5105 | 0.5914 | 0.7553 | 0.9597 | 0.6566 | 0.8718 | 0.9285 | 1.0000 | 5.5530 | 0.9633 | 0.6957 | 1.2680 | 1.6950 |
| 0.5362 | 0.3708 | 0.5076 | 0.5774 | 0.7321 | 0.5252 | 0.8696 | 0.6439 | 1.0000 | 0.9799 | 1.1470 | 1.4210 | 1.1751 |
| 0.3181 | 0.7067 | 0.7220 | 0.3776 | 0.6200 | 0.8606 | 1.0340 | 1.0380 | 1.0210 | 1.0000 | 0.9447 | 1.1700 | 1.4540 |
| 0.2237 | 0.3262 | 0.6002 | 0.4792 | 0.5163 | 0.3994 | 0.5141 | 1.4370 | 0.8718 | 1.0590 | 1.0000 | 0.9362 | 1.1010 |
| 0.4658 | 0.2538 | 0.4155 | 0.3750 | 0.5653 | 0.5118 | 0.6242 | 0.7886 | 0.7037 | 0.8547 | 1.0680 | 1.0000 | 0.9769 |
| 0.2024 | 0.2467 | 0.3372 | 0.6568 | 0.5132 | 0.3431 | 0.5417 | 0.5901 | 0.8497 | 0.6876 | 0.9082 | 1.0240 | 1.0000 |

* cigen value $=13.343$
* vector $=$
$\begin{array}{llllllllllllll}1.00000 & 0.88277 & 0.71639 & 0.70760 & 0.60639 & 0.62915 & 0.49856 & 0.46627 & 0.38128 & 0.41478 & 0.33556 & 0.31258 & 0.27811\end{array}$

Table $1113 \times 13 X==\left[x_{1}\right]$ (usual AHP)

| 1 | 2 | 3 | 4 | 5 | 6 | 7 | 8 | 9 | 10 | 11 | 12 | 13 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| 1 | 1.774 | 1.071 | 2.175 | 1.919 | 0.5758 | 1.045 | 12.48 | 16.16 | 3.908 | 0.8363 | 1.902 | 3.884 |
| 0.5636 | 1 | 0.8771 | 3.227 | 1.781 | 0.4639 | 0.2651 | 1.411 | 1.370 | 0.9083 | 1.924 | 8.971 | 3.072 |
| 0.9338 | 1.140 | 1 | 6.522 | 1.000 | 2.094 | 0.4587 | 4.016 | 3.730 | 1.012 | 3.680 | 0.6331 | 2.824 |
| 0.4599 | 0.3099 | 0.1533 | 1 | 7.611 | 0.9521 | 1.959 | 0.1990 | 0.2546 | 0.8926 | 2.651 | 0.5842 | 3.456 |
| 0.5212 | 0.5615 | 0.999 | 0.1314 | 1 | 1.357 | 1.393 | 1.208 | 0.8352 | 1.392 | 8.979 | 6.493 | 0.7722 |
| 1.737 | 2.156 | 0.4776 | 1.050 | 0.7372 | 1 | 0.6666 | 2.710 | 8.762 | 10.76 | 0.6958 | 1.390 | 3.436 |
| 0.9574 | 3.772 | 2.180 | 0.5104 | 0.7178 | 1.500 | 1 | 1.085 | 0.7848 | 2.004 | 14.61 | 6.686 | 3.485 |
| 0.08013 | 0.7087 | 0.2490 | 5.024 | 0.8277 | 0.3691 | 0.921 | 1 | 1.584 | 0.5654 | 6.730 | 1.110 | 7.981 |
| 0.06187 | 0.7300 | 0.2681 | 3.927 | 1.197 | 0.1141 | 1.274 | 0.6312 | 1 | 1.145 | 1.349 | 0.5351 | 2.335 |
| 0.2559 | 1.101 | 0.9879 | 1.120 | 0.7182 | 0.09297 | 0.4989 | 1.869 | 0.8730 | 1 | 2.956 | 0.8129 | 1.162 |
| 1.196 | 0.5197 | 0.2718 | 0.3772 | 0.114 | 1.437 | 0.06842 | 0.1486 | 0.7415 | 0.3383 | 1 | 0.8814 | 1.807 |
| 0.5258 | 0.1115 | 1.579 | 1.712 | 0.1540 | 0.7196 | 0.1496 | 0.9006 | 1.869 | 1.230 | 1.135 | 1 | 0.05509 |
| 0.2574 | 0.3255 | 0.3541 | 0.2894 | 1.295 | 0.2911 | 0.2869 | 0.1253 | 0.4283 | 0.8563 | 0.5535 | 18.15 | 1 |

* cigen value $=19.80$
* eigen vector $=$
$\begin{array}{llllllllllll}1 & 0.4795 & 0.6415 & 0.4551 & 0.4517 & 0.6925 & 0.7016 & 0.50 \%) & 0.3069 & 0.2703 & 0.2202 & 0.2479\end{array} 0.3613$

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At the end of this section we note that we can use LLS estimation ( $\rightarrow \S 3$ ) for our purpose.
First we apply LLS estimation $(\rightarrow(3.8))$ to Table 6 (Example 3) and have $\hat{w}_{i}^{\prime}$ as estimate of $w_{i}(i=1, \cdots, 7)$ (in Table 4). These are shown in Table 12, where $\hat{w}_{i}$ (in Table 6) are shown again for the comparison. (Of course $\hat{w}_{i}^{\prime} \mathrm{s}$ are standardized as $\hat{w}_{1}^{\prime}=1$ ). We can see that $\hat{w}_{i}^{\prime}$ gives a surprisingly good approximation to $\hat{w}_{i}(i=1, \cdots, 7)$.

Applying LLS again to Table 7 we have Table 13. This also gives fairly good approximations. Next we have Table 14 from Table 10 and Table 15 from Table 11 by the same way as above.

| Table 12 <br> (Decomposition) |  |  |
| :--- | :--- | :--- |
| $i$ | $w_{i}$ | $w_{i}^{\prime}$ |
| 1 | 1 | 1 |
| 2 | 0.6067 | 0.6061 |
| 3 | 0.0498 | 0.0498 |
| 4 | 0.1222 | 0.1218 |
| 5 | 0.1136 | 0.1137 |
| 6 | 0.2916 | 0.2904 |
| 7 | 0.2131 | 0.2134 |
|  | (eigen | $(1.1 S)$ |
|  | vector) |  |


| Table 13 |  |  |
| :---: | :---: | :---: |
| (AHP) |  |  |
| $i$ | $\omega_{i}$ | $\omega_{i}{ }^{\prime}$ |
| 1 | 1 | 1 |
| 2 | 0.7679 | 0.78601 |
| 3 | 0.0793 | 0.0755 |
| 4 | 0.1303 | 0.1273 |
| 5 | 0.1460 | 0.1385 |
| 6 | 0.4139 | 0.4067 |
| 7 | 0.2201 | 0.2224 |
|  | (rigen vector) | (1.1.5) |



Generally we can say that the LLS estimation gives a good approximation to the eigen vector estimation when the observation error is small and the number of objects is small.

The labor of the calculation of the LLS is far easier than the eigen vector method. The former is easily done on a desk calculator of pocket size, but the latter needs at least a personal computer. So the advantage of LLS method should be highly appreciated even in the highly computerized countries like Japan, USA and etc. Nevertheless the author thinks that the deep meaning of AHP is concealed in using the eigen vector for a estimation.

## Conclusions

For the statistical model stated in (3.1) and (3.2) our BIBD method gives better results than the usual AHP method. Further the LLS method (3.8) gives very good approximation to the eigen value method. In our present research we neglect the block effects (that is, (4.3) holds commonly for any blocks). But the evaluation criterion often depends on the judge himself. So for such cases we should modify the assumption that (3.1) or (4.3) hold.

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