

Seemingly Unrelated Regression Model with I(d) Regressors ($d > 1/2$) and Its Estimation

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I. Introduction

We derive the asymptotic properties of GLSE of Seemingly Unrelated Regression Model with nonstationary I(d) regressors ($d > 1/2$).

II. The Model and Assumptions

Let us consider the following seemingly unrelated regression model with nonstationary I(d) regressors x_{1t} and x_{2t} , where L is a lag operator.

$$y_t = \alpha_t + \beta_t x_t + u_t, \\ (1-L)^d x_t = w_t, \quad w_t = \psi_t(L) \varepsilon_t, \quad i=1,2, \quad t=1,2,\dots,T.$$

We make the following Assumptions 1 and 2.

Assumption 1 $d_i > \frac{1}{2}$, $\psi_t(L) = \sum_{j=0}^{\infty} \psi_{tj} L^j$ ($\psi_{t0} = 1$), $\sum_{j=0}^{\infty} j |\psi_{tj}| < \infty$, and all roots of $\psi_t(z) = 0$ are outside the unit circle. $x_{it} = 0$ ($t \leq 0$). ($i=1, 2$)

Assumption 2 $\begin{pmatrix} \varepsilon_{1t} \\ \varepsilon_{2t} \end{pmatrix} \sim IID \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \begin{pmatrix} \sigma_{\varepsilon_{11}} & 0 \\ 0 & \sigma_{\varepsilon_{22}} \end{pmatrix} \right)$, $\begin{pmatrix} u_{1t} \\ u_{2t} \end{pmatrix} \sim IID \left(\begin{pmatrix} 0 \\ 0 \end{pmatrix}, \Sigma \right)$, $\Sigma = \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix}$.
 ε_{it} and u_{jt} ($i, j=1,2, \quad t,s=1,2,\dots,T$) are independent.

The model can be written in conventional matrix form as

$$y = X\theta + u,$$

where y, X, u and θ are given as follows :

$$y = (y_1', y_2')', \quad y_1 = (y_{11}, y_{12}, \dots, y_{1T})', \quad y_2 = (y_{21}, y_{22}, \dots, y_{2T})',$$

$$X = \begin{pmatrix} e & x_1 & 0 & 0 \\ 0 & 0 & e & x_2 \end{pmatrix}, \quad e = (1, 1, \dots, 1)', \quad x_1 = (x_{11}, x_{12}, \dots, x_{1T})', \quad x_2 = (x_{21}, x_{22}, \dots, x_{2T})',$$

$$u = (u_1', u_2')', \quad u_1 = (u_{11}, u_{12}, \dots, u_{1T})', \quad u_2 = (u_{21}, u_{22}, \dots, u_{2T})'$$

$$\theta = (\theta_1', \theta_2')', \quad \theta_1 = (\alpha, \beta)', \quad \theta_2 = (\alpha_2, \beta_2)'$$

III. Estimation of θ and Its Properties

Let $\hat{\Sigma}$ be the estimator of Σ given below.

$$\hat{\Sigma} = \begin{pmatrix} \hat{\sigma}_{11} & \hat{\sigma}_{12} \\ \hat{\sigma}_{21} & \hat{\sigma}_{22} \end{pmatrix}$$

where $\hat{\sigma}_y = \frac{1}{T^{-2}} \tilde{u}_i' \tilde{u}_j$, and \tilde{u}_i is the OLS residual vector.

Then the (feasible) GLSE of θ is defined as follows,

$$\hat{\theta} = (X'(\hat{\Sigma} \otimes I_T)^{-1} X)^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} y$$

Using the normalizer given by

$$D = \text{Diag}(T^{\frac{1}{2}}, T^{d_1}, T^{\frac{1}{2}}, T^{d_2}),$$

$D(\hat{\theta} - \theta)$ can be written as

$$D(\hat{\theta} - \theta) = (D^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} X D^{-1})^{-1} D^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} u.$$

We can show that $D^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} X D^{-1}$ and $D^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} u$ converge weakly to the following.

$$D^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} X D^{-1} \Rightarrow \begin{pmatrix} \sigma_{22} H_{11} & -\sigma_{12} H_{12} \\ -\sigma_{12} H_{21} & \sigma_{11} H_{22} \end{pmatrix} \equiv H$$

where $H_{11} = \begin{pmatrix} 1 & \int_0^1 F_{d_1-1}(r) dr \\ * & \int_0^1 F_{d_1-1}^2(r) dr \end{pmatrix}$, $H_{12} = \begin{pmatrix} 1 & \int_0^1 F_{d_2-1}(r) dr \\ \int_0^1 F_{d_1-1}(r) dr & \int_0^1 F_{d_1-1}(r) F_{d_2-1}(r) dr \end{pmatrix}$, $H_{21} = H_{12}'$, $H_{22} = \begin{pmatrix} 1 & \int_0^1 F_{d_2-1}(r) dr \\ * & \int_0^1 F_{d_2-1}^2(r) dr \end{pmatrix}$,

$F_{d_i-1}(r) \equiv (\psi_i(1) \sigma_{\epsilon_{ii}}^{\frac{1}{2}} / \Gamma(d_i)) \int_0^r (r-s)^{d_i-1} dW_{\epsilon_i}(s)$ and $W_{\epsilon_i}(s)$ is standard Brownian motion.

$$D^{-1} X'(\hat{\Sigma} \otimes I_T)^{-1} u \Rightarrow \begin{pmatrix} \sigma_{22} W_{u_1}(1) - \sigma_{12} W_{u_2}(1) \\ \sigma_{22} \int_0^1 F_{d_1-1}(r) dW_{u_1}(r) - \sigma_{12} \int_0^1 F_{d_1-1}(r) dW_{u_2}(r) \\ -\sigma_{12} W_{u_1}(1) + \sigma_{11} W_{u_2}(1) \\ -\sigma_{12} \int_0^1 F_{d_1-1}(r) dW_{u_1}(r) + \sigma_{11} \int_0^1 F_{d_2-1}(r) dW_{u_2}(r) \end{pmatrix} \equiv K.$$

Therefore we get the following Theorem.

Theorem

$$D(\hat{\theta} - \theta) \Rightarrow H^{-1} K.$$

References

- Billingsley, P. (1968), *Convergence of Probability Measure*, John Wiley, New York.
 Maekawa, K. and H. Hisamatsu. (1996), SUR Models with I(1) Regressors.